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1    **Association of the Functional Movement Screen™ with match-injury burden in men's**  
2    **community rugby union.**

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4

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13

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15

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27 **Ethics approval** Research Ethics Approval Committee for Health, University of Bath.

28 **Abstract**

29 Evidence supporting the use of Functional Movement Screen (FMS™) to identify athletes'  
30 risk of injury is equivocal. Furthermore, few studies account for exposure to risk during  
31 analysis. This study investigated the association of FMS™ performance with incidence and  
32 burden of match-injuries in adult community rugby players. 277 players performed the  
33 FMS™ during pre-season and in-season time-loss injuries and match exposure were  
34 recorded. The associations between FMS™ score, pain, and movement-pattern  
35 asymmetries with match-injury incidence ( $\geq 8$  days time-loss/1000 hours), severe match-  
36 injury incidence ( $> 28$  days time-loss/1000 hours), and match-injury burden (total time-loss  
37 days/1000 hours for  $\geq 8$  days match-injuries) were analysed using Poisson regression.  
38 Multivariate analysis indicated players with pain and movement-pattern asymmetry  
39 during pre-season had 2.9 times higher severe match-injury incidence (RR, 90%CI=2.9,  
40 0.9-9.7) and match-injury burden (RR, 90%CI=2.9, 1.3–6.6). Players with a typically low  
41 FMS™ score (mean – 1SD threshold) were estimated to have a 50% greater match-injury  
42 burden compared to players with a typically high FMS™ score (mean + 1SD threshold) as  
43 match-injury burden was 10% lower per 1-unit increase in FMS™ score. As the strongest  
44 association with injury outcome was found for players with pain and asymmetry, when  
45 implementing the FMS™ it is advisable to prioritise these players for further assessment  
46 and subsequent treatment.

## 47    **Introduction**

48    In men's community rugby union, one player receives an injury causing them to miss at  
49    least one game every three team games (Roberts, Trewartha, England, Shaddick, &  
50    Stokes, 2013). On average, each of these injuries requires 7.6 weeks out of competition in  
51    order to recover (Roberts et al., 2013). However, injury risk factors in men's community  
52    rugby are poorly understood with the exception of previous injury, which has consistently  
53    been identified as a risk factor for further injury (Chalmers, Samaranayaka, Gulliver, &  
54    McNoe, 2012; Quarrie et al., 2001). As such, more information is needed to inform injury  
55    reduction strategies.

56    One approach to understanding the likelihood of a player getting injured is to conduct  
57    screening. However, comprehensive screening such as the medical screening protocol  
58    developed for the Australian College of Sports Physicians (Brukner, White, Shawdon and  
59    Holzer, 2004) can be too costly, too time consuming and may require practitioner  
60    expertise that is not available within community clubs. A simple and quick-to-perform  
61    movement control assessment has the potential to be of great benefit to community  
62    teams. Compared with comprehensive athlete screening protocols, the Functional  
63    Movement Screen™ (FMS™) is more economical to administer and can be performed by  
64    individuals with basic FMS™ training (Cook, Burton, & Hoogenboom, 2006a, 2006b). The  
65    FMS™ comprises seven movement patterns that assess individuals' strength, balance and  
66    range of motion and are combined with three movements that screen for pain (Cook et  
67    al., 2006a, 2006b). The primary function of the FMS™ is to identify areas of movement  
68    deficiency in individuals, but it has also been used to predict injury in a range of athletic  
69    populations, with conflicting results concerning the relationship of the FMS™ scores with  
70    injury. The FMS™ was not associated with injury in runners (Hotta et al., 2015), mixed

71 sports (including cross-country, football, soccer, swimming, tennis, and volleyball) high  
72 school athletes (Bardenett et al., 2015), mixed sports (including basketball, football,  
73 volleyball, track and Field, swimming, soccer, golf and tennis) NCAA division 1 athletes  
74 (Warren, Smith, & Chimera, 2015), or professional soccer players (Zalai, Panics, Bobak,  
75 Csaki, & Hamar, 2015). However, associations of FMS™ with injury have been identified in  
76 collision based sports, including American football (Kiesel, Butler, & Plisky, 2014; Kiesel,  
77 Plisky, & Voight, 2007) and rugby union (Duke, Martin, & Gaul, 2017; Tee, Klingbiel,  
78 Collins, Lambert, & Coopoo, 2016). In American Football, FMS™ score (Kiesel et al., 2007)  
79 and presence of movement-pattern asymmetry (Kiesel et al., 2014) were associated with  
80 a higher likelihood of injury. In elite rugby union, movement competency (Duke et al.,  
81 2017; Tee et al., 2016) and sub-test scores (Tee et al., 2016) were associated with  
82 increased likelihood of injury, but movement-pattern asymmetry and likelihood of injury  
83 were poorly associated (Duke, et al., 2017).

84 One of the most important risk factors for rugby injury is the amount of time players are  
85 exposed to risk (Williams et al., 2017) yet no study described above accounted for  
86 exposure. Only a few sports-based FMS™ studies have accounted for players' exposure  
87 during analysis (Chalmers et al., 2018; Chalmers et al., 2017; Hammes, Aus der Fünter,  
88 Bizzini, & Meyer, 2016). In veteran football players, Hammes et al. (2016) reported no  
89 clear association between FMS™ score and playing time until first injury. In junior  
90 Australian Football players, Chalmers et al. (2017) also reported no association between  
91 FMS™ score and injury. However, the presence of one or more asymmetries was  
92 associated with 1.9 times higher likelihood of injury in junior Australian Football players,  
93 escalating to 2.8 times likelihood of injury where players had 2 or more asymmetries  
94 (Chalmers et al., 2017). Following a direct replication of the Australian Football study

95 design, the results originally presented in 2017 could not be replicated, and asymmetry  
96 during FMS™ testing was not associated with a significant increase in prospective injury  
97 in the replication dataset (Chalmers et al., 2018). As such, asymmetry should be  
98 considered when analysing the association between FMS™ performance and rugby injury.

99 This study investigated FMS™ performance (including the influence of movement  
100 asymmetry and pain), while accounting for individual player match exposure, the  
101 association with time-loss match-injury outcomes of 8 days or greater, and what FMS™  
102 score was associated with the greatest difference in match-injury burden for a men's  
103 community rugby population.

104

## 105 **Methods.**

106 This study was designed as a prospective observational cohort study. All participants  
107 performed the FMS™ at the beginning of the study period after which match-injury and  
108 exposure data were collected over a competitive rugby season.

### 109 *Participants*

110 Participants were recruited from the community rugby playing population in England. A  
111 similar population has previously been categorised into three sub-groups as Semi-  
112 professional (Rugby Football Union (RFU) levels 3-4; highest level of English community  
113 rugby), Amateur (RFU levels 5-6) and Recreational (RFU levels 7-9) (Roberts et al., 2013).  
114 An inclusion criteria was that participating clubs had to have a recognised qualified sports  
115 therapist, osteopath, chiropractor, physiotherapist, or doctor to record injuries. At the  
116 time of recruitment, participants were injury free (self-reported) and all were considered  
117 by the coaching team to be eligible and under consideration to play in the club's 1<sup>st</sup> team

118 for the forthcoming season. In total, 23 clubs (men's senior squad only) were recruited  
119 (Figure 1), from which 433 players volunteered to participate.

120

121 \*\*\*FIGURE 1 NEAR HERE\*\*\*

122

### 123 *Ethical approval and consent*

124 Participating clubs were provided with study information and full instructions for testing  
125 procedures prior to the testing session taking place, which was then disseminated to all  
126 players who provided written informed consent at the start of the testing session. Ethics  
127 approval was granted by the University of Bath, Research Ethics Approval Committee for  
128 Health (EP 12/13 58).

### 129 *Examiners*

130 Fourteen people acted as raters during the testing period, attending participating clubs in  
131 groups of 4. All raters had a sports science background and included undergraduate  
132 students, post graduate students, and academic staff. Rater training was received from a  
133 certified FMS™ trainer and five of the raters had over 12-months experience using FMS™  
134 prior to this study. No formal reliability study was performed as part of the present study,  
135 though raters with similar and varied backgrounds have previously been shown to have  
136 good intra-rater (interclass correlation coefficient (ICC), 95% confidence interval (CI), =  
137 0.81, 0.69-0.92) and inter-rater reliability (ICC, 95% CI, = 0.81, 0.70-0.92) when delivering  
138 the FMS™ (Bonazza, Smuin, Onks, Silvis, & Dhawan, 2017).



139     *Procedures*

140     FMS™ data were collected during pre-season (between July and September 2013) at each  
141     club. After an introduction to the testing procedures by the research team leader,  
142     participants signed informed consent forms. Participants' self-reported primary playing  
143     position and age (years) and the research team recorded height (m) (Leicester Height  
144     Measure, Seca, UK) and mass (kg) (SC-240 body composition monitor, Tanita, USA).  
145     Participants' movement control, pain and movement pattern asymmetry were then  
146     assessed using the FMS™ in an indoor area within the club.

147     *Functional Movement Screen™*

148     Participants wore shorts, T-shirts, their normal trainers and were divided into four even  
149     groups with one researcher completing the entire FMS™ screen with each group.  
150     Participants were not allowed to complete a warm-up or to perform preparatory  
151     stretching prior to testing. The FMS™ was conducted using the standard method (Cook et  
152     al., 2006a, 2006b). For each movement pattern component, a central demonstration with  
153     standard verbal instructions was provided by the research team leader to ensure that all  
154     participants received the same information prior to screening. Participants were not  
155     aware of the scoring system. Each component was repeated up to three times by  
156     participants and the best scores recorded. Component movement scores were recorded  
157     in real-time by the raters who were able to change their viewing position. FMS™  
158     components were scored on an ordinal scale (0-3), where 'zero' is given if the participant  
159     experiences pain during the test, through to a score of 'three' for perfect test execution.  
160     For bilateral movement patterns (inline lunge, rotational stability, shoulder mobility,  
161     active straight leg raise and hurdle step) scores were recorded for both right and left  
162     sides. Asymmetry was present if the movement scores for the left and right sides differed

163 by one point or more. Where a difference in score was recorded for a bilateral movement  
164 pattern, the lower score for was used when the overall FMS<sup>TM</sup> score was calculated. A  
165 player's FMS<sup>TM</sup> score was calculated according to standardised criteria (Cook et al., 2006a,  
166 2006b).

#### 167 *Match exposure*

168 For every 1<sup>st</sup> team match of the 2013-14 rugby season, participating clubs recorded  
169 individual player match exposure using a standardised form. Match exposure was  
170 recorded as 20, 40, 60 or 80 minutes.

#### 171 *Player injury*

172 Injury management staff at participating clubs completed and returned injury forms. Any  
173 injury incurred during a first team match resulting in an absence from participation in full  
174 training or match play for 8 days or more from the day of the injury was defined as a  
175 "time-loss" match-injury (Fuller et al., 2007). The date on which the injured player was fit  
176 for game selection (whether or not they actually played on that date) was recorded as the  
177 return to play date. Injury severity was calculated as the number of days elapsed between  
178 the date of injury and 'return to play' date.

179 For all time-loss injuries, information was recorded for the anatomical site, injury type,  
180 injury event, treatment, time of injury during match and severity using a standard report  
181 form. Injury diagnoses were recorded using the Orchard Sports Injury Classification  
182 System version 8 (Rae, Britt, Orchard, & Finch, 2005) by the injury management staff.  
183 Only injuries incurred during match play were recorded and therefore absences from  
184 match play due to illness or injuries incurred through any other activity (including rugby  
185 training) were excluded.

186 *Statistical Analysis*

187 Data analysis was performed using SPSS (Version 22 for Windows, Armonk, NY. IMB  
188 Corp). Descriptive characteristics for player demographics were reported as mean  $\pm$   
189 standard deviation (SD). Mean FMS™ scores were compared according to players' injury  
190 status ('injured' = any player suffering a time-loss injury during the season, or 'non-  
191 injured' = no time-loss injury during the season).

192 Injury incidence rates (IIRs) were reported per 1000 player match-hours and severity  
193 recorded as the number of days absence from full training or match play. Match-injury  
194 burden was reported as total time-lost (days) per 1000 player match-hours. The sum of  
195 match-injuries and sum of total match exposure was used to calculate incidence of overall  
196 ( $\geq 8$  days time-loss) and severe match-injuries ( $> 28$  days time-loss). Effect sizes (ES) were  
197 quantified and considered as trivial ( $\leq 0.2$ ), small ( $> 0.2-0.6$ ), moderate ( $> 0.6-1.2$ ), large  
198 ( $> 1.2-2.0$ ) and very large ( $> 2.0-4.0$ ) (Batterham and Hopkins, 2006). A General Estimating  
199 Equation (GEE) was used to determine associations between FMS™ score, asymmetry,  
200 pain and injury count. Multivariate analyses were undertaken and over-dispersion was  
201 controlled for using a Pearson chi-square scaling parameter (McCullagh and Nedler,  
202 1989). Regression analysis was offset for exposure (hours) and was adjusted for club  
203 (cluster), playing level stratification (semi-professional; amateur; recreational) and player  
204 (random effects). Analysis was performed for any match-injury ( $\geq 8$  days time-loss), severe  
205 match-injury ( $> 28$  days time-loss) and match-injury burden (time-lost days) for all  $\geq 8$  days  
206 time-loss injuries. Results are presented as rate ratio (RR) with 90% confidence intervals  
207 (90%CI) and interpreted using clinical-magnitude based inference (Hopkins and  
208 Batterham, 2016). Threshold values for unlikely/harmful (25) and most/very unlikely (5)

209 were used to derive the odds ratio for making mechanical inference (Hopkins and  
210 Batterham, 2016).

211

## 212 **Results**

### 213 *Descriptive summary*

214 Due to factors including club withdrawal from the study, individual players never playing  
215 for the 1<sup>st</sup> team or otherwise returning incomplete data, time-loss injury and individual  
216 match exposure data were reported for 277 (64%) of the initial 433 players who were  
217 screened. For the 277 players included within the analysis, FMS™ and anthropometric  
218 characteristics are presented in table 1.

219

220 \*\*\*TABLE 1 NEAR HERE\*\*\*

221

222 For the 277 players the median FMS™ score was 14 (mean ± standard deviation (SD) =  
223 14.1±2.6), 28% of all players reported pain and 72% of all players displayed asymmetry on  
224 ≥1 of the FMS™ movement patterns. Twenty-three percent of all players displayed both  
225 movement-pattern asymmetry and reported pain, while 23% of all players displayed  
226 neither asymmetry nor reported pain when completing FMS™ screening. Both  
227 movement-pattern asymmetry and pain were most commonly reported for the shoulder  
228 mobility movement pattern.

229

230

231 Of the 277 players, 57 (21%) players sustained 74 acute match-injuries across 4359 player  
232 match-hours (equivalent to 218 team-games) (Table 2). No recurrent or gradual onset  
233 injuries were reported. Overall match-injury incidence ( $\geq 8$  days time-loss) was 17.0  
234 (90%CI=14.0–20.6) injuries/1000 player match-hours. Of the 57 injured players, 30  
235 players accumulated 35 severe ( $>28$  days time-loss) match-injuries with an incidence of  
236 8.0 (90%CI=6.1–10.6) severe match-injuries/1000 player match-hours. For all  $\geq 8$  days  
237 time-loss match-injuries the match-injury burden was 655 (90%CI=541-792) days/1000  
238 player match-hours. Contact ( $n = 57$ ) and non-contact injuries ( $n = 9$ ) accounted for 77%  
239 and 12% of match-injuries, respectively, while no event was reported for 8 (11%) match-  
240 injuries.

241

242 *\*\*\*TABLE 2 NEAR HERE\*\*\**

243

244 The greatest match-injury burden was associated with injuries involving the knee (127.3  
245 days/1000 player match-hours), ankle (84.2 days/1000 player match-hours) and the  
246 shoulder (70.7 days/1000 player match-hours; table 3), while the match-injury types  
247 associated with the greatest match-injury burden were ligament tears/sprains (163.6  
248 days/1000 player match-hours), muscle tears/strains (92.0 days/1000 player match-  
249 hours) and fractures (76.6 days/1000 player match hours; table 4).

250

251 *\*\*\*TABLE 3 NEAR HERE\*\*\**

252 \*\*\*TABLE 4 NEAR HERE\*\*\*

253

254 *Association of FMS™ score with injury outcomes*

255 The distribution of FMS™ scores for these 277 players, stratified by injury status is  
256 displayed in Figure 2. Difference in mean FMS™ score between players with any match-  
257 injury ( $14.0 \pm 2.7$ ) and non-injured players ( $14.1 \pm 2.6$ ) was trivial (Figure 2; Effect size  
258 (ES), 90% CI= -0.04, -0.27–0.19). The difference in mean FMS™ score between players  
259 who sustained a severe match-injury ( $13.5 \pm 2.6$ ) and non-injured players ( $14.1 \pm 2.6$ ) was  
260 also trivial (Figure 2; ES, 90% CI= -0.22, -0.53 – 0.09).

261

262 \*\*\*FIGURE 2 NEAR HERE\*\*\*

263

264 Poisson regression analysis indicated the association of FMS™ score and injury incidence  
265 was trivial for overall match-injury (RR, 90%CI=0.96, 0.90-1.02) and severe match-injury  
266 (RR, 90%CI=0.92, 0.84-1.01) (Figure 4). A 1-unit increase in FMS™ score was associated  
267 with a possibly beneficial 10% lower match-injury burden (RR, 90%CI=0.90, 0.83-0.97).  
268 Rate ratio analysis was used to determine the FMS™ score associated with the greatest  
269 difference in match-injury burden (Figure 3). Players scoring  $\geq 16$  (31%) compared with  
270  $< 16$  on the FMS™ demonstrated the greatest difference in all match-injury outcomes  
271 including a very likely beneficial 59% lower match-injury burden (RR, 90%CI=0.41, 0.22-  
272 0.76), a likely beneficial 51% lower severe match-injury incidence (RR, 90%CI=0.49, 0.24-

273 1.02) and a likely beneficial 30% lower overall match-injury incidence (RR, 90%CI=0.70,  
274 0.47-1.05).

275

276 \*\*\*FIGURE 3 NEAR HERE\*\*\*

277

#### 278 *Association of pain and asymmetry with injury*

279 Multivariate Poisson regression analysis indicated that the presence of any movement  
280 pattern asymmetry was associated with a very likely harmful 2.5 times higher severe  
281 match-injury incidence (RR, 90%CI=2.5, 1.0–6.2) and very likely harmful 2.4 times higher  
282 match-injury burden (RR, 90%CI=2.4, 1.4–4.3) (Figure 4) compared with players with no  
283 movement pattern asymmetry, adjusted for FMS<sup>TM</sup> score. The presence of pain was  
284 associated with a likely harmful 1.8 times higher match-injury burden (RR, 90%CI = 1.8,  
285 1.0–3.2) compared with players who did not report pain during movement pattern  
286 testing, adjusted for FMS<sup>TM</sup> score.

287

288 \*\*\*FIGURE 4 NEAR HERE\*\*\*

289

290 Players displaying asymmetry without pain (n=136, 49%) were associated with a likely  
291 harmful 2.3 times higher incidence of severe match-injury (RR, 90%CI=2.3, 0.8-6.5) and  
292 likely harmful 2.2 times higher match-injury burden (RR, 90%CI=2.2, 1.1-4.4) compared  
293 with the control group (Figure 5), adjusted for FMS<sup>TM</sup> score. Players presenting both  
294 asymmetry and pain (n=65, 23%) were associated with a likely harmful 2.9 times higher

295 incidence of severe match-injury (RR, 90%CI=2.9, 0.9-9.7) and a very likely harmful 2.9  
296 times higher match-injury burden (RR, 90%CI=2.9, 1.3–6.6) compared with the control  
297 group, adjusted for FMS™ score.

298

\*\*\*FIGURE 5 NEAR HERE\*\*\*

299

## 300 **Discussion**

301 This study investigated whether the Functional Movement Screen™ score, pain and/or  
302 asymmetry determined prospectively during FMS™ testing were associated with time-  
303 loss match-injury outcomes in men's community rugby players. Better movement control,  
304 indicated by a higher FMS™ score, was associated with less time lost to injury, where a 1-  
305 point increase in FMS™ score was associated with a 10% lower match-injury burden.  
306 Controlling for FMS™ score, the presence of both pain and movement asymmetry were  
307 associated with an approximately 3-fold increase in severe match-injury incidence and  
308 match-injury burden. While players with an FMS™ score of  $\geq 13$  demonstrated a clearly  
309 beneficial lower match-injury burden compared to players scoring  $< 13$ , the greatest  
310 difference in all injury outcomes was found for players scoring  $\geq 16$  compares to players  
311 scoring  $< 16$ .

312

313 This study was the first to investigate FMS™ and injury burden and used Poisson linear  
314 regression offset for player match exposure to analyse players risk of injury. As a measure  
315 of movement competency, a 1-point increase in FMS™ performance was associated with



316 a 10% lower injury burden, which implies that players with better movement patterns  
317 lose less time to injury than players with deficient movement patterns. However, no  
318 meaningful association between FMS<sup>TM</sup> score and overall match-injury incidence ( $\geq 8$ -days  
319 time-loss) or severe match-injury ( $> 28$ -days time-loss) was found. The lack of association  
320 between FMS<sup>TM</sup> score and match-injury incidence may be due to the many random  
321 events and player to player contacts that occur during rugby match play, which makes  
322 predicting 'who' gets injured challenging. Previous researchers have likened the ability of  
323 the FMS<sup>TM</sup> to predict 'who' will get injured to flipping a coin (Dorrel, Long, Shaffer and  
324 Myer, 2018). Yet better movement competency was associated with lower match-injury  
325 burden for which there is no clear and obvious rationale. A possible explanation is that  
326 players with better movement competency (higher FMS<sup>TM</sup> scores) are able to achieve and  
327 better maintain 'optimal' body positions during contact events such as the tackle, ruck  
328 and maul compared with players with poor movement competency (lower FMS<sup>TM</sup> scores).  
329 For example, improved lower-limb alignment during a tackle situation may reduce forced  
330 knee valgus when under the sudden external load experienced by the tackler, resulting in  
331 a lower match-injury burden. Hopkins, Marshall, Batterham, & Hanin (2009) recommend  
332 making inferences by comparing the effect of different levels of continuous predictors  
333 i.e., comparing the injury burden for players with typically low (mean-SD) to typically high  
334 (mean+SD) scores. In this study, the mean FMS<sup>TM</sup> score for all players was 14.1 (SD = 2.6).  
335 A 2SD improvement in players' FMS<sup>TM</sup> score thus approximates to 50% lower match-  
336 injury burden based on this relationship. A similar result was reported for veteran soccer  
337 players where players with a 'low' FMS<sup>TM</sup> score (FMS<sup>TM</sup>  $< 10$ ) had 1.9 times the injury  
338 incidence compared to those with an 'intermediate' FMS<sup>TM</sup> score (FMS<sup>TM</sup> = 10-14)  
339 (Hammes et al., 2016). These results support the notion that better movement

340 competency (higher FMS<sup>TM</sup> score) is associated with lower injury outcomes. As FMS<sup>TM</sup>  
341 scores have been demonstrated to be modifiable by implementing movement control  
342 interventions (Kiesel et al., 2011), clubs may be advised to maximise players movement  
343 competency by intervention post screening. Improving players movement competency  
344 should be considered by clubs as even moderate reductions in injury burden may have  
345 worthwhile effects on competition outcomes (Williams et al., 2016).

346

347 In the present study, the presence of  $\geq 1$  asymmetry was associated with 2.2 times the  
348 overall injury burden (664 vs 291 days/1000 player match-hours) and 2.3 times the  
349 incidence of severe injury (8.6 vs 3.7 injuries/1000 player match-hours) when adjusted for  
350 FMS<sup>TM</sup> score. When assessing sports injury risk, recommended methods of analysis  
351 include Cox regression, frailty modelling (Finch and Marshall, 2016) and linear regression  
352 (Bahr and Holme, 2003) where the forms of analysis account for individual player  
353 exposure to the risk (participation in the sport). While the present study used Poisson  
354 linear regression, two previous studies of contact sports have used Cox regression in their  
355 research of FMS<sup>TM</sup> and injury outcome. In Australian Rules Football, junior players with  $\geq 1$   
356 movement asymmetry were associated with 1.9 times the likelihood of injury (any trauma  
357 or medical condition resulting in match time-loss) compared with players with no  
358 asymmetry, which increased to 2.8 times the likelihood of injury for players with  $\geq 2$   
359 movement pattern asymmetries (Chalmers et al., 2017). In addition, players that  
360 displayed both pain and asymmetry had a 1.6 times likelihood of time-loss injury  
361 (Chalmers et al., 2017). However, these results have not yet proven to be replicable in  
362 junior Australian Rules Football (Chalmers et al., 2018). In the present study, players that

363 demonstrated both pain and asymmetry had a likely harmful 2.9 times higher incidence  
364 of severe injury and very likely harmful 2.9 times higher injury burden for players  
365 displaying both pain and asymmetry when adjusted for FMS™ score. What is not  
366 apparent when conducting the FMS™ is why asymmetry or pain is present. Possible  
367 reasons could be related to hand and leg dominance, poor training practice or previous  
368 injury. Clubs using the FMS™ may be advised to triage players displaying asymmetry or  
369 pain for further investigation by a registered medical practitioner, such as a  
370 physiotherapist, to identify the underlying cause, for which a corrective exercise  
371 programme may be developed. Priority for such referral should be granted to players who  
372 display asymmetry and also report pain as these players were associated with a greater  
373 risk of injury than asymmetry alone.

374

375 Most sports screening tests measure using a continuous scale and must be translated to a  
376 dichotomous outcome (Bahr, 2017). In the present study, rate ratio analysis was used to  
377 determine whether a FMS™ score would maximise the difference in injury outcomes.  
378 Players (31%) that scored  $\geq 16$  on the FMS™ had beneficially lower injury outcomes,  
379 including overall injury incidence (12.4 v 18.9 injuries / 1000 player match hours), severe  
380 injury incidence (4.6 v 9.5 injuries / 1000 player match-hours) and injury burden (325 v  
381 794 days / 1000 player match-hours) compared to players scoring  $< 16$ . Similar scores  
382 have been proposed by studies in different populations including intercollegiate athletics  
383 (FMS™  $\leq 17$ ; Weise, Boone, Mattacola, McKeon and Uhl, 2014), physically active students  
384 (FMS™  $< 17$ ; Letafatkar, Hadadnezhad, Shojaedin and Mohammadi, 2014) and National  
385 Collegiate Athletic Association Division II athletes (FMS  $\leq 15$ ; Dorrel et al., 2018). However,

386 a score of  $\geq 16$  contrasts with other FMS<sup>TM</sup> literature where a score of FMS<sup>TM</sup>  $\leq 14$  has  
387 commonly been proposed as an injury predictive value (Kiesel et al., 2007; Chorba et al.,  
388 2010; Butler et al., 2013; Lisman et al., 2013). These previous studies did not account for  
389 participants' exposure when identifying their injury predictive values using receiver  
390 operator characteristic analysis (Keisel et al., 2007, Butler et al., 2013) and otherwise  
391 adopted the cut-off score of FMS<sup>TM</sup>  $\leq 14$  based on previous research (Chorba et al., 2010;  
392 Lisman et al., 2013). While a score of  $\geq 16$  is higher than the commonly proposed score of  
393  $> 14$ , no previous literature has considered injury burden, used Poisson regression  
394 analysis, nor accounted for players match exposure with similar resolution, which likely  
395 effected these results. Overall, the better a player's movement competency, the lower  
396 the overall injury risk where a target score of FMS<sup>TM</sup>  $\geq 16$  should be employed to maximise  
397 the injury risk benefit.

398

399 No study has measured players' FMS<sup>TM</sup> scores and used the results to produce an exercise  
400 intervention demonstrated to be effective in reducing the injury risk of athletes. Many  
401 variables affect FMS<sup>TM</sup> scores which are player specific, possibly requiring an individualised  
402 approach to each player's pre-habilitation intervention. The FMS<sup>TM</sup> total score does not  
403 represent a unidimensional construct (Kazman, Galecki, Lisman, Deuster and O'Connor,  
404 2014), in that two players can have the same FMS<sup>TM</sup> score but achieve it with  
405 considerably different movement competencies. As such, a uniform solution to improve  
406 movement competency is not possible to prescribe based on total FMS<sup>TM</sup> score alone.  
407 During follow-up assessment of players highlighted as at 'higher risk', therapists must  
408 focus on the players specific movement deficiencies before providing a

409 treatment/intervention. Based on the proportion of players in the present study with low  
410 FMS™ scores, pain and/or asymmetry, if community club therapists started screening  
411 during pre-season, it is unlikely that the follow-up assessments necessary to determine  
412 each player's dysfunction and subsequent treatment would be complete until early into  
413 the competitive season, where the risk of injury is highest (Garraway and Macleod, 1995;  
414 Quarrie et al., 2001; Roberts et al., 2013). Rather than using FMS™ in isolation, clubs are  
415 advised to administer movement competency injury prevention programmes to all  
416 players during training, as such interventions have reduced injury in rugby (Attwood,  
417 Roberts, Trewartha, England, & Stokes, 2017; Hislop et al., 2017), football (Emery and  
418 Meeuwisse, 2010; Gilchrist et al., 2008; Soligard et al., 2010), basketball (Longo et al.,  
419 2012) and handball (Andersson, Bahr, Clarsen, & Myklebust, 2016; Olsen, Myklebust,  
420 Engebretsen, Holme, & Bahr, 2005). By implementing club wide movement control  
421 programmes such as Activate (Attwood, Roberts, Trewartha, England, & Stokes, 2017;  
422 Hislop et al., 2017) clubs would already be implementing a recommended player welfare  
423 strategy while adequate time is allocated to facilitate FMS™ screening and subsequent  
424 player follow-up to develop individualised programmes for 'higher risk' players. The  
425 implementation of Activate, FMS™ screening and subsequent player specific corrective  
426 treatment may have a combined and beneficial effect on player welfare and thus  
427 maximise the injury reduction benefit for limited resources available to community rugby  
428 teams.

429

430 *Strengths and Limitations of the Study*

431 Strengths of the study include the large sample of players followed throughout a season  
432 and the inclusion of individual players match exposure during analysis, as has been  
433 recommended when investigating injury risk factors (Bahr and Holme, 2003). This was  
434 also the first study to apply Poisson regression analysis, while accounting for playing level,  
435 which has previously been associated with significant differences in injury incidence  
436 (Roberts et al., 2013). There were some limitations to this study. Injury reporting was  
437 limited to match-injuries with a severity of  $\geq 8$ -days rather than 1-day. This injury  
438 definition excluded all training injuries and any match injuries  $< 8$ -days time-loss from the  
439 analysis, which do account for a small proportion of the overall injury burden. This  
440 approach was thought to be appropriate as it negated the need to report injury and  
441 exposure data for a squad of players at every training session, thus helping to maintain  
442 clubs' involvement in the study. As described in the methods, no formal reliability study  
443 was performed to determine agreement between assessors. The analysis performed  
444 throughout this study, was not powered for, and does not account for the type of injuries  
445 sustained which could influence the associations reported, due to the low count per  
446 injury type / site. As such, type and site of injury were limited to descriptive analysis only.  
447 Further investigation into the relationship between injury severity, injury burden, FMS<sup>TM</sup>  
448 score and specific injury types, such as anterior cruciate injury or hamstring injury as two  
449 examples, is recommended to affirm the association between the burden of specific  
450 injuries and movement competency screened using the FMS<sup>TM</sup>.

451  
452 Using the Functional Movement Screen<sup>TM</sup> to assess movement competency during pre-  
453 season may help practitioners to identify players at greater risk of match injury. Players  
454 movement competency should be maximised by practitioners, since a 1-point change in

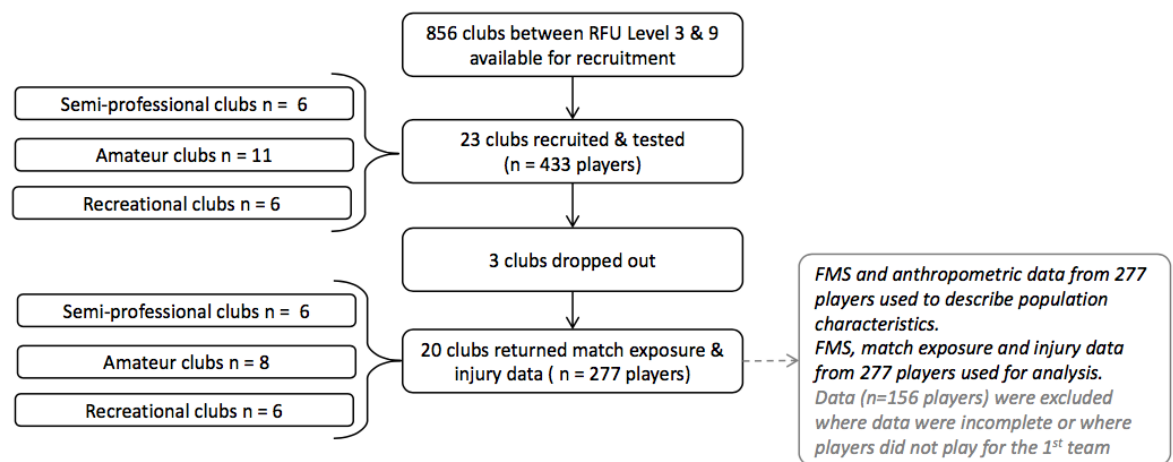
455 FMS™ score was associated with a 10% lower match-injury burden, resulting in a 50%  
456 lower match-injury burden when comparing players with typically low to typically high  
457 FMS™ scores. However, if screening started at the beginning of pre-season, some players  
458 may not receive corrective treatment until the early in-season period, due to the time  
459 required to conduct FMS™ screening, to follow-up and develop interventions for players  
460 identified as 'high risk'. As movement control programmes such as 'Activate' reduce  
461 rugby players injury burden, rugby clubs should implement Activate club-wide while  
462 screening is conducted in order to help maximise the welfare of their players. Following  
463 screening, players with the lowest FMS™ scores should be prioritised, particularly those  
464 with low FMS™ scores that report pain and display asymmetrical movements, as the  
465 combined presence of these factors was associated with the greatest injury risk.

466

467

## 468 Tables and Figures

469



470

471 Figure 1. Overview of the reach of the study, including the number of clubs that

472 participated, dropped-out, and volume of data used for analysis.

473

474 Table 1. FMS™ and anthropometric characteristics of 277 players, organised by playing  
475 level stratification.

Playing level	Clubs (n)	Players (n)	FMS score mean (SD)	Age (years) mean (SD)	Height (cm) mean (SD)	Mass (kg) mean (SD)	BMI (kg/m <sup>2</sup> ) mean (SD)
Semi-professional	6	85	14.2 (2.9)	23.7 (3.8)	182.2 (6.5)	95.7 (13.1)	28.8 (3.2)
Recreational	8	108	14.2 (2.4)	25.3 (4.2)	181.2 (6.6)	95.7 (13.1)	29.2 (3.7)
Amateur	6	84	13.8 (2.5)	25.4 (5.3)	178.7 (6.5)	91.2 (12.7)	28.6 (4.0)
Total	20	277	14.1 (2.6)	24.8 (4.4)	180.7 (6.5)	94.2 (13.0)	28.9 (3.6)

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477

478



Table 2. A summary of the nature and number of injuries including match-injury incidence and match-injury burden organised by playing level stratification.

	Injury count (n)	Contact injuries <u>n</u> (%)	Non-contact injuries <u>n</u> (%)	Unknown event injuries <u>n</u> (%)	Exposure (hours)	Incidence per 1000 player match-hours (90% CI)	Time lost (days)	Burden per 1000 player match-hours (90% CI)
Semi-professional	25	16 (64)	3 (12)	6 (24)	1272	19.7 (13.3-29.1)	781	614 (572-659)
Amateur	27	22 (81)	5 (19)	0	1683	16.0 (11.0-23.4)	1217	723 (684-765)
Recreational	22	15 (68)	6 (27)	1 (5)	1404	15.7 (10.3-23.8)	855	609 (570-651)

\*Percentages for contact and non-contact injuries do not sum to 100% where injury event details weren't fully reported.

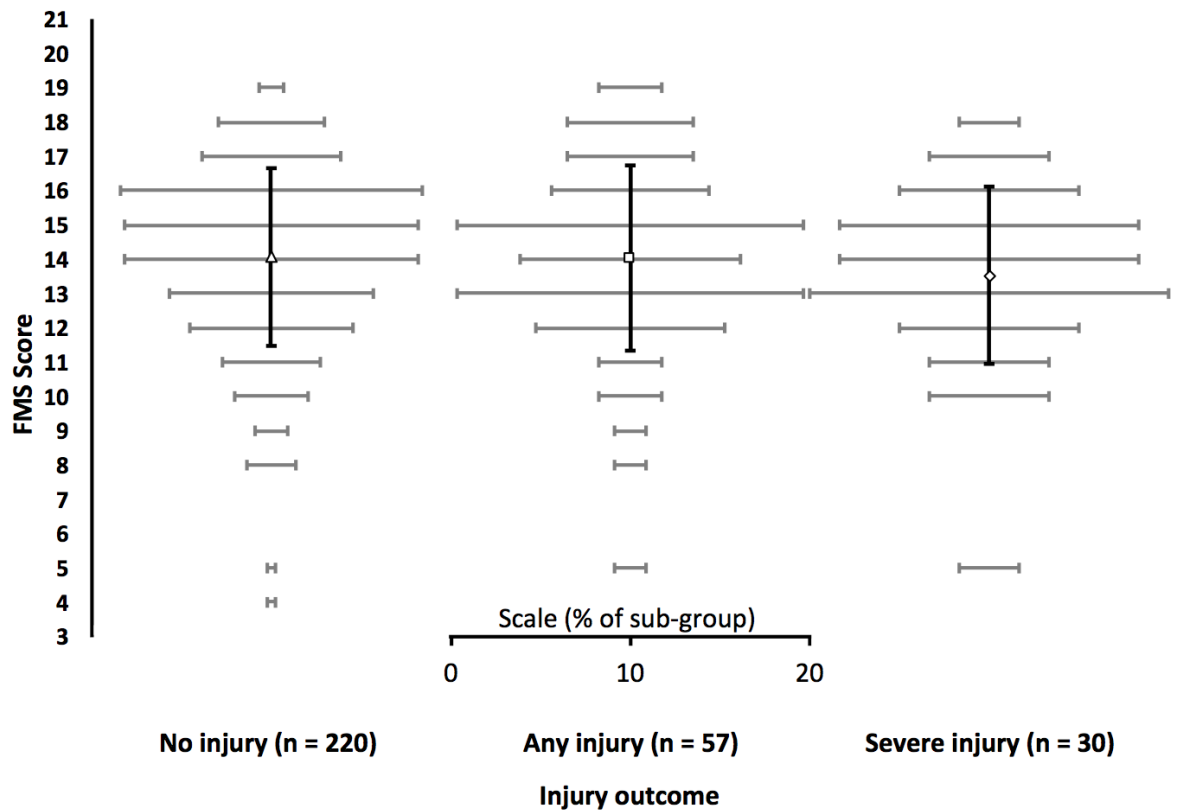
Table 3. The injury sites with greatest burden for all groups, arranged in descending order of match-injury burden.

	Injury count (n)	Contact injuries <u>n</u> (%)	Non-contact injuries <u>n</u> (%)	Unknown event injuries <u>n</u> (%)	Exposure (hours)	Incidence per 1000 player match-hours (90% CI)	Time lost (days)	Burden per 1000 player match-hours (90% CI)
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\*Percentages for contact and non-contact injuries do not sum to 100% where injury event details weren't fully reported.

Table 4. The injury types with greatest burden for all groups, arranged in descending order of match-injury burden.

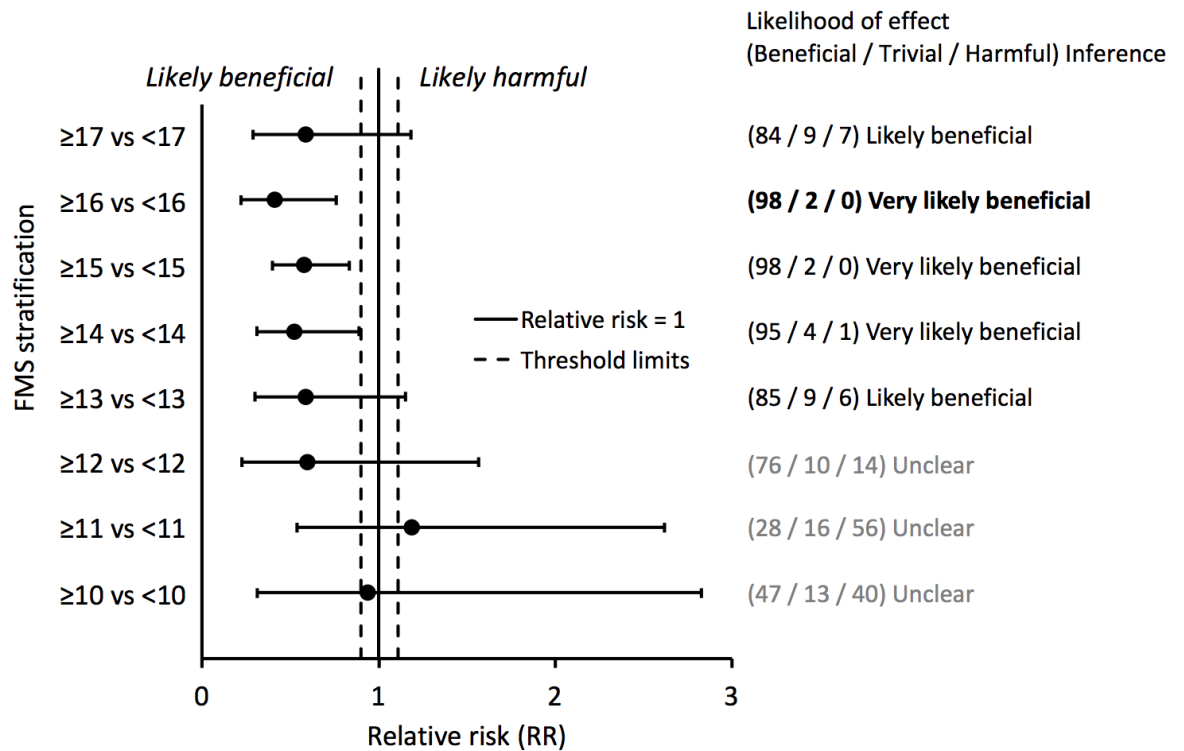
Injury type	n	Time-lost (days)	Exposure (hours)	Incidence per 1000 player match-hours (90% CI)	Mean severity days (90% CI)	Burden per 1000 player match-hours (90% CI)
Ligament tear/strain	15	713	4359	3.4 (2.3-5.3)	48 (31-73)	164 (107-250)
Muscle tear/sprain	15	401	4359	3.4 (2.3-5.3)	27 (18-41)	92 (60-140)
Fracture	7	334	4359	1.6 (0.9-3.0)	48 (26-89)	77 (41-142)
Nerve injury	7	217	4359	1.6 (0.9-3.0)	31 (17-58)	50 (27-93)
Tendon injury	5	212	4359	1.1 (0.6-2.4)	42 (20-88)	49 (23-101)



491

492 Figure 2. FMS™ scores stratified by injury definition; no injury, any injury ( $\geq 8$  days), and  
 493 severe injury ( $> 28$  days). Horizontal error bars represent frequency of FMS™ scores,  
 494 vertical error bars represent mean and 90% confidence limits.

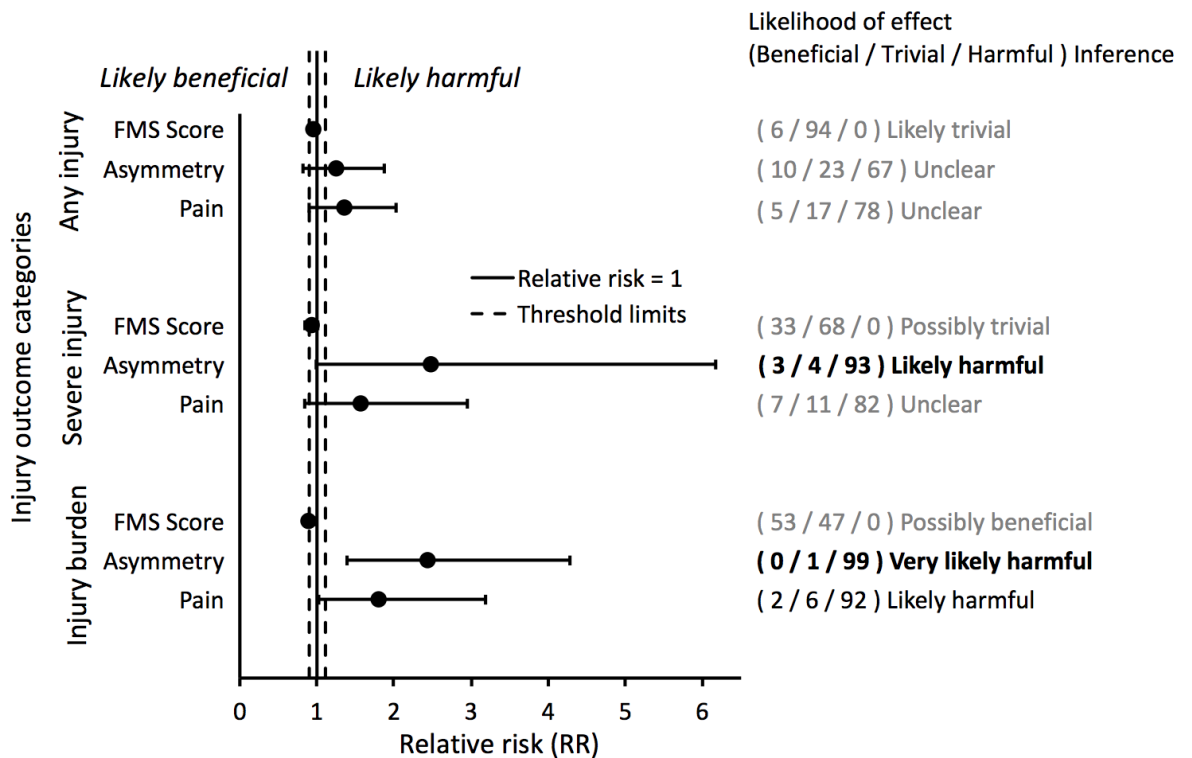
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497 Figure 3. Forest plot comparing match-injury burden (days/1000 player match-hours) by  
 498 FMS™ score stratification. The right side of the figure displays the likelihood of effect.  
 499 FMS™ scores at and above which resulted in a lower injury burden with a high likelihood  
 500 of effect are highlighted in bold (right column).

501

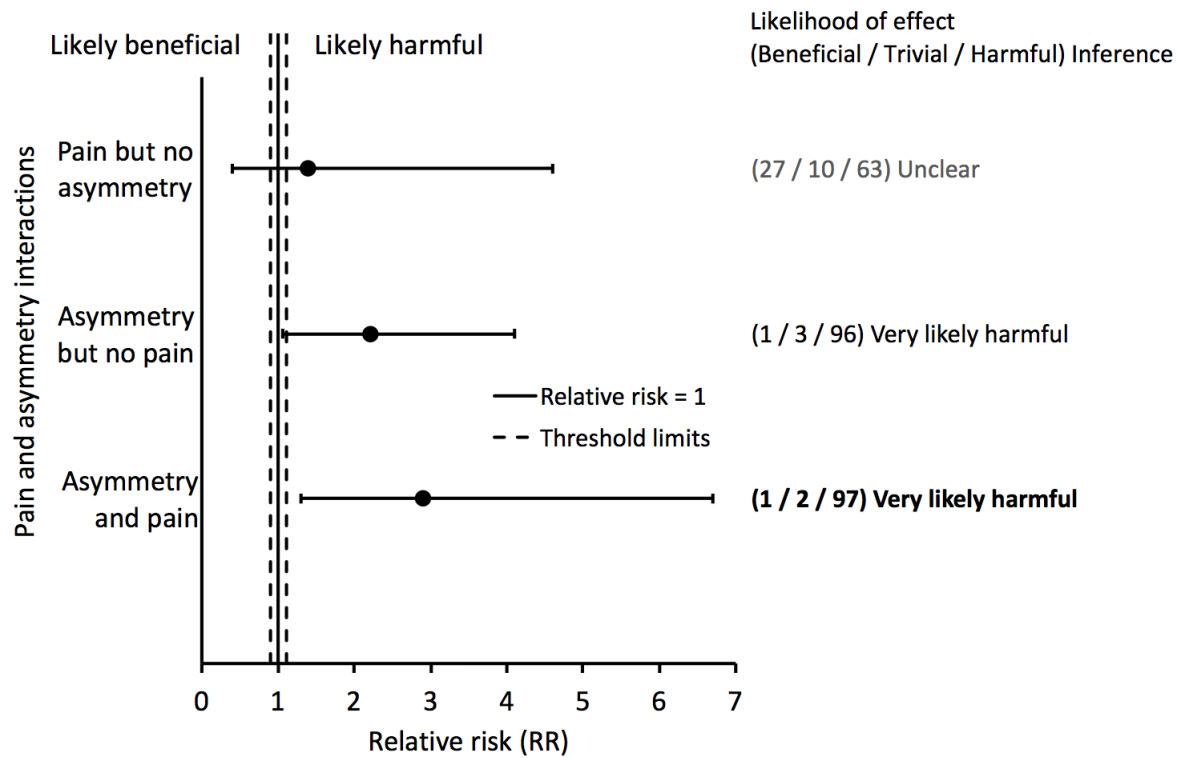


502

503 Figure 4. Forest plot displaying univariate results for relative risk of players with higher  
 504 FMS™ score (continuous) compared to lower FMS™ score; players displaying any  
 505 asymmetry compared to players with no asymmetry; and players reporting pain to  
 506 players not reporting pain. The largest effects are highlighted in bold.

507

508



509

510 Figure 5. Forest plot displaying the interaction effects of pain and asymmetry on match-

511 injury burden (days/1000 player match-hours) compared baseline (no asymmetry, no

512 pain). The largest effects are highlighted in bold.

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